

# Global Biogeochemical Cycles

## RESEARCH ARTICLE

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### Key Points:

- Multisource topsoil organic carbon prediction and prediction variance in Mexico and the conterminous United States
- Calculated stocks of 46–47 Pg of SOC (0- to 30-cm depth, years 1991–2010) using a simulated annealing regression framework
- Predicted stocks >30% below recent global estimates that are largely based on legacy data

### Supporting Information:

- Supporting Information S1
- Table S1
- Table S2

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## Soil Organic Carbon Across Mexico and the Conterminous United States (1991–2010)

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**Abstract** Soil organic carbon (SOC) information is fundamental for improving global carbon cycle modeling efforts, but discrepancies exist from country-to-global scales. We predicted the spatial distribution of SOC stocks (topsoil; 0–30 cm) and quantified modeling uncertainty across Mexico and the conterminous United States (CONUS). We used a multisource SOC dataset (>10 000 pedons, between 1991 and 2010) coupled with a simulated annealing regression framework that accounts for variable selection. Our model explained ~50% of SOC spatial variability (across 250-m grids). We analyzed model variance, and the residual variance of six conventional pedotransfer functions for estimating bulk density to calculate SOC stocks. Two independent datasets confirmed that the SOC stock for both countries represents between 46 and 47 Pg with a total modeling variance of  $\pm 12$  Pg. We report a residual variance of  $10.4 \pm 5.1$  Pg of SOC stocks calculated from six pedotransfer functions for soil bulk density. When reducing training data to define decades with relatively higher density of observations (1991–2000 and 2001–2010, respectively), model variance for predicted SOC stocks ranged between 41 and 55 Pg. We found nearly 42% of SOC across Mexico in forests and 24% in croplands, whereas 31% was found in forests and 28% in croplands across CONUS. Grasslands and shrublands stored 29 and 35% of SOC across Mexico and CONUS, respectively. We predicted SOC stocks >30% below recent global estimates that do not account for uncertainty and are based on legacy data. Our results provide insights for interpretation of estimates based on SOC legacy data and benchmarks for improving regional-to-global monitoring efforts.

## 1. Introduction

Terrestrial ecosystems store >1,500 Pg of soil organic carbon (SOC, approximate stock at 1-m depth) worldwide, but accurate spatial representation of these stocks is needed for fully understanding the contribution of soils within the global carbon cycle (Crowther et al. 2016, Food and Agriculture Organization [FAO] 2017). For global modeling and validating the SOC stored in terrestrial ecosystems, high-resolution gridded datasets such as the SoilGrids250m system (Hengl et al., 2017) are increasingly being used to describe spatial SOC patterns (Harden et al., 2017; Jackson et al., 2017) and trends (Naipal et al., 2018; Hengl & MacMillan 2019). Such datasets are also required to facilitate the formulation of reliable climate change adaptation guidelines and the establishment of regional to global carbon monitoring and information systems (Ciais et al., 2014; Stockmann et al., 2015; Vargas et al., 2017; Villarreal et al., 2018). Previous studies suggest that the greatest source of discrepancy across regional to global carbon cycling estimates is associated with the SOC pool (Crowther et al., 2016; Jones et al., 2005; Jones & Falloon, 2009; Murray-Tortarolo et al., 2016; Tifafi et al., 2017). Arguably, current scientific challenges associated with the discrepancy of the soil carbon pool are to quantify (1) the size and distribution of local to regional SOC stocks at scales relevant

to inform land management decisions (FAO, 2017, Banwart et al., 2017); (2) the amount of carbon losses from soils due to heterotrophic respiration (Bond-Lamberty et al., 2018), the amount of carbon removed from erosion (Naipal et al., 2018) or aquatic export (Tank et al., 2018), or changes in land use and land cover (Sanderman et al., 2017); and (3) the carbon emissions from impacts of past and future climate conditions (Crowther et al., 2016; Delgado-Baquerizo et al., 2017; Walsh et al., 2017). Solving these scientific challenges around carbon cycling requires a good understanding of different uncertainty sources around SOC datasets and SOC modeling efforts.

Major uncertainties in spatial SOC estimates that are extrapolated from points/pedons to continuous estimates across the land surface are related to several factors. These include measurement methods, data sources (SOC data and SOC environmental covariates) and their resolution and extent, the different periods of data collection, or using multiple modeling and evaluation strategies (Grunwald, 2009; Ogle et al., 2010; Stockmann et al., 2013). Thus, there is a need for improving interoperability for compiling the best available information and describing SOC spatial variability across local to global scales (Vargas et al., 2017).

Global modeling outputs for SOC represent the only estimates of SOC across large areas of the world without in situ ground observations. These global estimates rely on large datasets that combine multiple SOC data collection periods and methods for calculating SOC stocks. These inconsistencies represent a known but unquantified bias for calculating SOC stocks (Poeplau et al., 2017). Thus, there is a need to test different modeling approaches across areas with high density of SOC observations to improve the accuracy, detail, and reliability of global SOC estimates (Vitharana et al., 2019).

The Harmonized World Soil Data Base (HWSD) or the harmonized soil property values for broadscale modeling (WISE30sec, Batjes, 2016) are probably the most commonly used datasets for spatially quantifying SOC stocks and its spatial variability patterns at the global scale (Köchy et al., 2015; O'Rourke et al., 2015). The HWSD provides the most complete global soil description from synthesizing many regional or national soil maps, but it uses a polygon-based approach that has intrinsic quality limitations such as coarse scale (e.g., >1-km pixels), discrepancy between national datasets and broad categorical generalizations (Folberth et al., 2016; Stoorvogel et al., 2016). Regional to global efforts to improve the spatial representation of the global SOC pool also include those by the International Union of Soil Sciences (Arrouays et al., 2017), the International Soil Resource Information Centre (ISRIC, e.g., Hengl et al., 2014; Batjes et al., 2017; Hengl et al., 2017), the International Soil Carbon Network (ISCN, Harden et al., 2017), the Land GIS project (<https://landgis.openeo.org>), and the GlobalSoilMap consortium (Arrouays et al., 2014; Sanchez et al., 2009). Another initiative is the recent call from the United Nations Food and Agricultural Organization (FAO) requesting the development of country specific frameworks for reporting continuous and spatially explicit SOC stocks and patterns (FAO, 2017). These efforts have contributed information for global estimates, along with methodologies useful for applying standardized protocols for harmonizing SOC measurements from multiple sources for SOC assessments. However, validating global SOC estimates and developing country- or region-specific (e.g., North America) SOC prediction frameworks are still needed for increasing knowledge by quantifying uncertainties while explaining the discrepancy of current SOC estimates (e.g., country specific to global scales).

Large discrepancies have been reported among global (Tifafi et al., 2017) or country-specific SOC estimates (Guevara et al., 2018). Consequently, reporting uncertainty and bias of SOC estimates will allow better parameterization of land surface models, improved local to regional monitoring baselines, and informed policy and management decisions regarding SOC stocks (Viscarra-Rossel et al., 2014). The current discrepancies around SOC estimates could be partially attributed to SOC sampling errors and bias in the SOC sampling locations, but this is information that may not be always available for improving SOC estimates. Other sources of errors and spatial artifacts are related with the use of different measurement methods (or analytical techniques) for quantifying SOC stocks, lack of information on bulk density (BD), or rock fragments, and different methods to apply pedotransfer functions may generate contrasting results. For predictive SOC mapping (McBratney et al., 2003), the quality of SOC training data and the quality of SOC environmental covariates represent a potential source of uncertainty that will propagate to final predictions. Thus, increasing information about how and when SOC data are collected and selecting only the most effective SOC environmental covariates (i.e., from remote sensing, geomorphometry, climatology surfaces, and thematic maps) will reduce the propagation of errors on further modeling efforts. Quantifying the errors

from inputs and models that influence SOC predictions and identifying how they are spatially distributed will benefit planning for future SOC sampling strategies, by assuming that a larger sample is required across areas with higher discrepancies and modeling bias (FAO, 2017; Heuvelink, 2014). Optimizing soil sampling strategies is constantly required to validate/calibrate SOC predictions and reduce their uncertainties across unsampled areas.

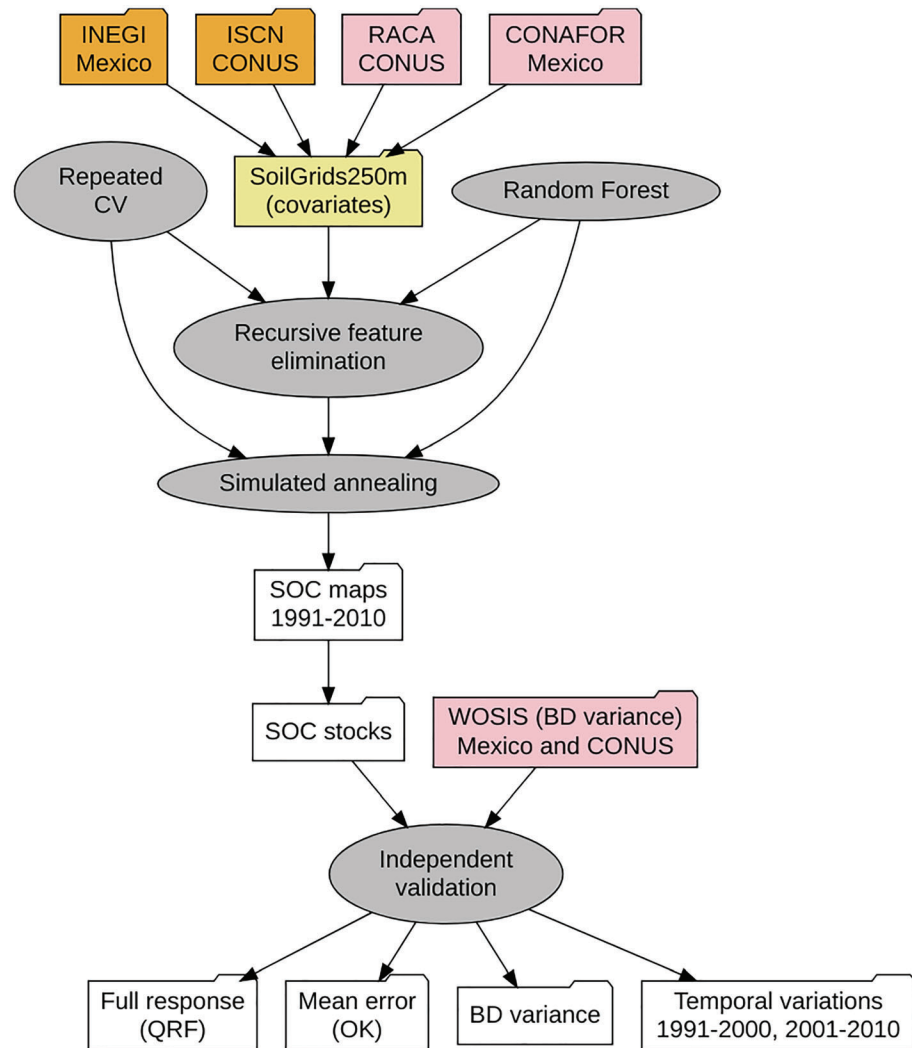
North America is a region characterized by a long history of soil data collection that has produced unprecedented information of SOC. For example, SOC predictions and estimates across Mexico and the conterminous United States (CONUS) are available from a variety of methods and in different formats. These include soil type polygon maps, field observations, and reflectance spectroscopy analysis, as well as global SOC variability gridded surfaces based on environmental correlation methods (e.g., Bliss et al., 2014; Hengl et al., 2014; Hengl et al., 2017; Wijewardane et al., 2016). Further examples include the use of linear geostatistics for the interpolation of SOC across Mexico (Cruz-Cárdenas et al., 2014) and SOC modeling efforts across the United States (Padarian et al., 2015). For increasing prediction accuracy of SOC models, flexible statistics (e.g., machine learning) have been proposed to better predict nonlinear relationships between SOC observational data and their environmental predictors at global and continental scales (Hengl et al., 2017; Ramcharan et al., 2018). Thus, SOC environmental covariates (i.e., surrogates of climate, biota, topography, and geology) and observational data can be coupled with machine learning algorithms to improve the representation of spatial variability and uncertainty in SOC stocks. Reducing uncertainties from different sources of information, increasing data-model agreement, and simplifying model complexity (by assessing variable importance and removing not informative SOC environmental covariates) are required to enable the fine-scale monitoring of SOC stocks across countries and regions of the world where no such information is otherwise available (de Gruijter et al., 2016; Minasny et al., 2017; Viscarra-Rossel et al., 2014).

In this study, we quantified the spatial variability and associated uncertainty of SOC stocks across different land use categories of CONUS and Mexico, two countries with rich information of SOC measurements. Previous analyses have shown large discrepancies in SOC stocks (0–30 cm, ranging from ~38 to ~92 Pg of SOC) derived from country-specific or global SOC estimates (Bliss et al., 2014; Hengl et al., 2017; Lajtha et al., 2018; Paz Pellat et al., 2016; Wieder et al. 2015). Our main goal was to generate a spatial predictive model of SOC variability for the top 30-cm depth at 250-m spatial resolution across both countries with information collected between 1991 and 2010.

We asked the following interrelated questions: (1) Which are the best SOC environmental covariates (i.e., prediction factors increasing SOC modeling accuracy) across Mexico and CONUS? (2) How much variation in SOC can machine learning methods (e.g., tree-based, kernel based, and probabilistic-based) explain across this region using repeated cross-validation? (3) What is the SOC variance associated with multiple calculation methods for estimating SOC stocks and what is the variance associated to multiple model predictions? (4) What are the sensitivities of these predictions associated to different training datasets (i.e., decadal information from different collection periods; 1991–2000 and 2001–2010)? The value of considering different collection periods is to explore the sensitivity of model predictions associated to different training datasets and provide insights for better interpretation of decadal changes in SOC stocks. In summary, this study provides benchmark information about how SOC spatial distribution is constrained by the soil forming environment (i.e., climate, biota, topography, and geology) and quantifies the variance (spatial and temporal) from using multiple SOC observational datasets.

## 2. Datasets and Methods

We followed a digital soil mapping strategy (Figure 1) for the prediction of the spatial variability of SOC across both countries. Digital soil maps are generated using field and laboratory observational methods coupled with environmental data through quantitative relationships (McBratney et al., 2003; Minasny et al., 2008, Minasny et al., 2013). We assumed that the spatial variability of SOC (represented by observational data) can be predicted across large geographical (unsampled) areas as a function of soil forming factors (climate, biota, topography, and geology; Jenny, 1941). These factors (surrogates of the soil forming environment) are represented through three main sources of information: remote sensing sensors, gridded climatology products (e.g., precipitation and temperature), and digital terrain analysis (i.e., geomorphometry; see Pike et al., 2009 and Wilson, 2012, McBratney et al., 2003).



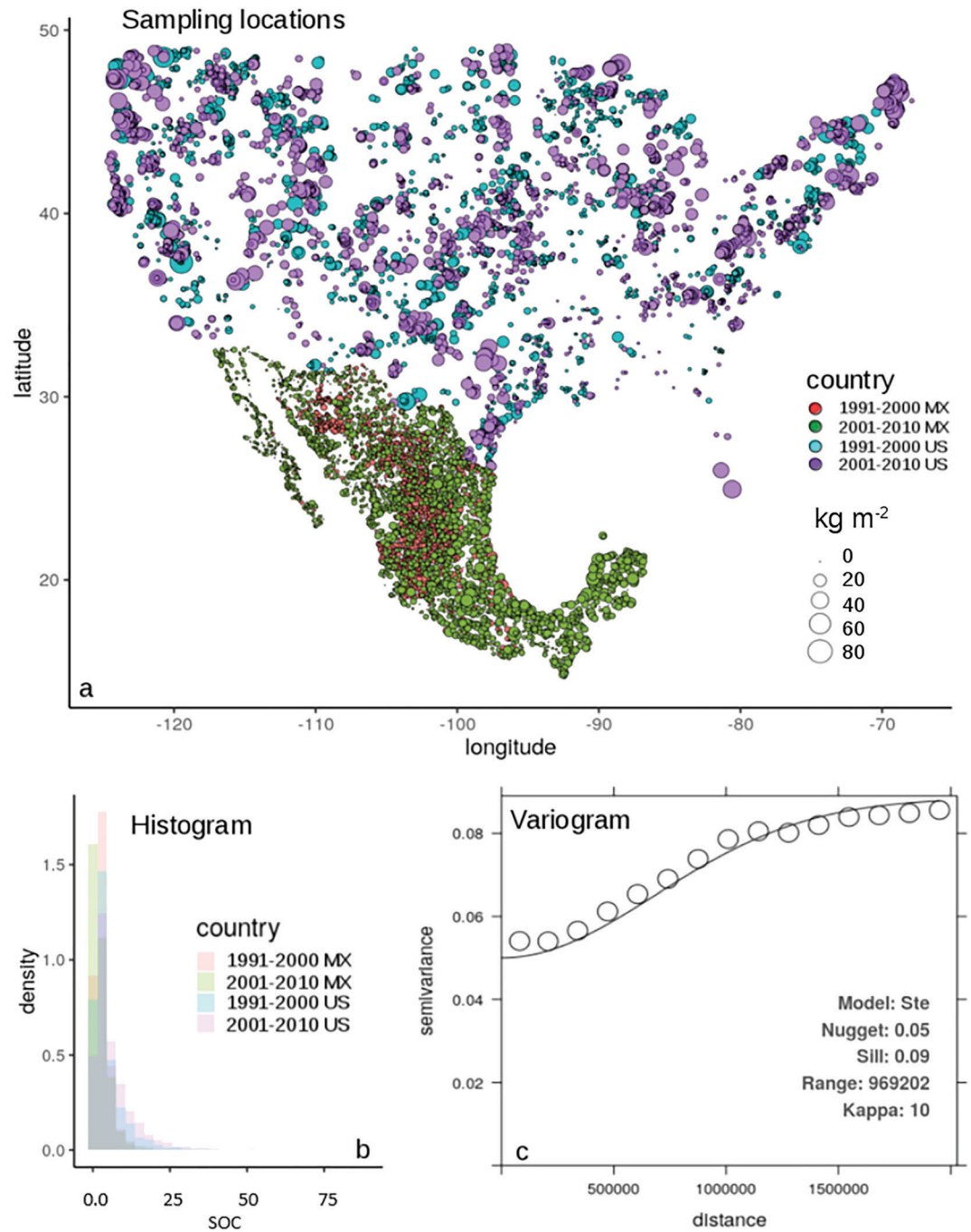
**Figure 1.** Flow diagram of the proposed methodology to predict the spatial variability of soil organic carbon (SOC) stocks across Mexico and conterminous United States (CONUS). The orange folders indicate the SOC data sources used for training, and the pink folders indicate SOC data sources used for validating. These sources were harmonized with the SoilGrids250m covariates. The white folders indicate the main results of this methodology. The gray ovals indicate main methodological steps. CV: five fold cross validation, QRF: quantile regression forest, OK: Ordinary Kriging, BD: bulk density.

## 2.1. SOC Observational Data

Legacy SOC estimates across CONUS were obtained from the ISCN (ISCN latest version 2018, >18,000 pedons available; Harden et al., 2017). Data from Mexico were provided by the Instituto Nacional de Estadística y Geografía (INEGI, SERIES 1, and 2;  $n > 65,000$  pedons available; Krasilnikov et al., 2013). We used only the observations collected between 1991 and 2010 to minimize confounding factors (related to potential changes in the SOC pool;  $n = 10,385$ ; Figure 2). We considered all soil horizons containing upper and lower soil depth limit information. The combination of using soil depth continuous functions (Bishop et al., 1999; Malone et al., 2009) and deriving the weighted average (by depth) from the first sampled soil horizon at 0-cm depth to all soil horizons sampled within the first 30 cm of soil depth, allowed aggregating irregular soil horizons for calculating SOC stocks across both countries. The weights for calculating these stocks were selected defining the proportion of each horizon within this 0- to 30-cm interval of soil depth.

Most contributors (across CONUS) and INEGI (in Mexico) considered (or adapted) the U.S. Department of Agriculture (USDA) Soil Taxonomy guidelines for interpreting soil surveys including SOC and other soil





**Figure 2.** Distribution and descriptive statistics of available datasets. (a) The point map shows the spatial sampling locations of available data for the period 1991–2010 (1991–2000 and 2001–2010). (b) The colored histograms are representing the statistical distribution of all datasets (i.e., combined conterminous United States and Mexico information). (c) The variogram (relation between distance and variance of observed values) and variogram parameters (nugget, sill, and range) are representing the spatial structure captured with available data. *Ste* is a Stein model parameterization (and its associated Kappa value) for the covariance function between all pairs of points separated by distance units (range in meters) defining the spatial structure of soil organic carbon available datasets.

variables (Soil Survey Staff., 1999). For CONUS, the ISCN database provides a harmonized compilation from many contributors (e.g., Natural Resources Conservation Service, U.S. Geological Survey, and site-specific research or academic groups; Harden et al., 2017). However, the largest contributor for this curated

dataset is the U.S. USDA Natural Resource Conservation Service, where the SOC concentration was mainly obtained by the Walkley-Black technique (Soil Survey Staff, 2014). All samples for Mexico were systematically collected and analyzed by INEGI (INEGI, 2014, Krasilnikov et al., 2013), and SOC concentration was also measured using the Walkley-Black technique. Potential error propagation from the use of different methods to calculate SOC using information collected over long periods of time (before 1991) is beyond the scope of this study. We only considered the sensitivity of SOC models (i.e., model outputs) to variations in training data and inputs derived from different pedotransfer functions for estimating BD (section 2.6).

## 2.2. Calculation of SOC Stocks

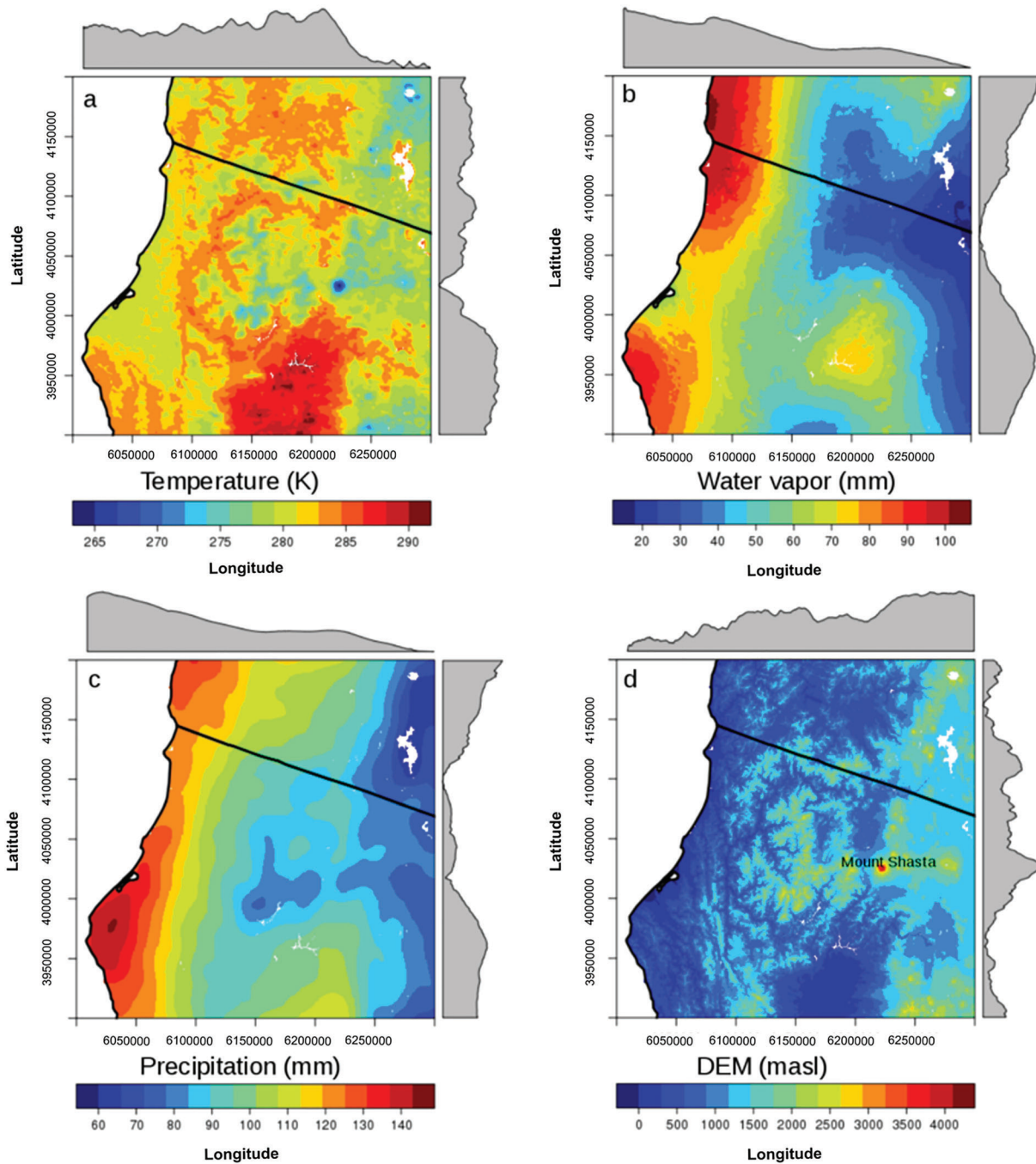
SOC stocks were derived by a linear combination of soil depth (0–30 cm), coarse fragment (CF) data, SOC concentration (%), and soil BD following the method proposed by Nelson and Sommers (1982) as implemented by Hengl et al. (2017). In general, CF across CONUS was measured in the field considering soil fragments >2 mm and direct gravimetric mass-methods. In Mexico, CF was also measured in field (considering soil fragments >2 mm) but expressed as percentage of gravel, stones, and pebbles. For the CONUS dataset, the ISCN has calculated SOC stocks using both modeled (i.e., incomplete) and nonmodeled (i.e., complete) information about the aforementioned variables (Harden et al., 2017). We only used information flagged as “complete” by the ISCN, so no model or pedotransfer functions were used for estimating BD and consequently SOC stocks. In Mexico, BD was estimated in the field using soil type maps, soil texture, soil organic matter, and soil structure following international soil mapping guidelines (FAO, 2006, p 51, Table 58). These guidelines are based in a rule-based approach originally described in the German soil-mapping guidelines (Ad-hoc-AG Boden, 2005) and have been applied to the collection and analysis of soils across Mexico (Siebe et al., 2006) and for the contribution of Mexico to the United Nations SOC map (FAO and ITPS, 2018).

## 2.3. The Environmental Covariate Space

For spatially representing soil forming factors (Jenny, 1941) we used environmental covariates from the SoilGrids250m system (Hengl et al., 2017). This dataset represents over 150 variables of environmental gridded data including terrain derivatives from a digital elevation model (DEM), the enhanced vegetation index (EVI), climate (precipitation and land surface temperature), and other soil-related gridded variables (Figure 3 and Table S1 in the supporting information). This covariate space is representative for the analyzed period of time (1991–2010) and is described in detail by previous publications (Hengl et al., 2017; Reuter & Hengl, 2012). We extracted this global information within the geographical limits of the NALCMS (North American Land Change Monitoring System, 77% of land use classification accuracy) at 250-m spatial resolution (Colditz et al., 2014).

## 2.4. Recursive Feature Elimination

We first performed a variable reduction strategy using of a recursive feature elimination technique (Kuhn, 2008). and multiple models were fitted repeatedly using all possible combinations of highly ranked predictors. Predictors were ranked using as indicator the cross-validated prediction error of a Random Forest tree ensemble. We selected the Random Forest as our overall accuracy indicator method because it showed the highest predictive capacity compared against different machine learning algorithms tested in our modeling selection strategy (Figure S1 in the supporting information). This method is based on bagging predictors and the combination of multiples regression trees derived from different random data subsets (Breiman, 2001). Each model grows with the number of trees for minimizing the prediction variance. Model parameters to define the number of predictors and subsets on each regression tree were automatically selected by the means of tenfold cross validation (Figure 1). Cross validation is a resampling technique that we used for maximizing the accuracy of results while obtaining a robust and stable prediction error estimate used for further selecting the most informative predictors. In addition, Random Forests uses an out-of-bag cross-validation form for assessing the relevance of each predictor in the model. Thus, multiple lists of the “best” predictors are generated from each Random Forest model realization in the recursive feature elimination framework. This provides a probabilistic assessment to determine the best predictors to retain at the end of the algorithm (Kuhn 2008). After a 5-times repeated fivefold cross validation for the recursive feature elimination technique (to account for the model sensitivity to data variations and reduce overfitting), we selected the first 25 environmental covariates for SOC.



**Figure 3.** Visualization of covariates across the political boundaries between California and Oregon in western conterminous United States. (a) Land surface temperature, (b) precipitable water vapor, (c) precipitation, and a digital elevation model; see Table S1 for detailed description and sources of these variables. The gray histograms represent variation across latitude or longitude.

### 2.5. Simulated Annealing

The 25 environmental covariates selected from the recursive feature elimination analysis were used on a simulated annealing regression framework for predicting SOC stocks. Simulated annealing is a well-known optimization framework for soil sampling designs (Minasny & McBratney, 2006; Szatmári et al., 2015) and for validating digital soil maps (Biwas & Zhang, 2018). Simulated annealing is a framework from

statistical mechanics and combinatorial optimization problems (Kirkpatrick, et al., 1983) that here we apply for maximizing the feature selection and prediction accuracy of SOC relevant environmental covariates.

In a simulated annealing framework, used for prediction (i.e., regression), a global search is performed and random perturbations are induced to the dataset for identifying the variables that are more sensitive to data variations and that have higher prediction capacity for the target variable (i.e., SOC). We used the cross-validated Random Forest error as indicator to analyze the effect of such perturbations. This process is constantly repeated, and many iterations are produced in a global learning search that should in theory result in better solutions (Kuhn 2008). We used the Random Forest regression algorithm within the simulated annealing framework to improve the probability of detecting the main drivers of SOC spatial patterns. After a 5-times repeated fivefold cross validation, the entire data set is used for generating a model in the last execution of the simulated annealing global search. This model is built on the predictor subset that is associated with the optimal number of iterations determined by the cross-validation resampling technique (Kuhn 2008). We used the final model of the simulated annealing framework for making predictions across 250-m grids reporting the first five ranked environmental covariates of each generated model. These environmental covariates were selected because they contributed the most to reducing the error in the global search of the simulated annealing iterative (i.e., tree ensemble learning) process.

## 2.6. Uncertainty Analysis

We represented uncertainty of our modeling approach as the sensitivity of prediction models to multiple data inputs. We first explored the residual variance of our SOC training data against six SOC stocks calculated using six BD pedotransfer functions. Then, we analyzed the spatial structure of these residuals using Geostatistics, and computed a model residual error against fully independent SOC datasets. Finally, we computed the modeling SOC prediction variance and the full quantile response of residuals (from independent datasets and from the BD variance) to the highest ranked environmental covariates (Figure 1). We postulate that estimates of SOC fall within a range of errors, and it is therefore important to account for variation in model inputs and model outputs. Our main goal was to quantify the variability range around predicted SOC stocks using multiple uncertainty indicators.

### 2.6.1. Pedotransfer Functions for Bulk Density Variance

We predicted SOC stocks at the pedon locations available in the WoSIS system (Batjes et al., 2017). We calculated the residual variance of our predictions and independent SOC stocks. These independent stocks were calculated using the WoSIS SOC concentration data (%), and six conventional pedotransfer functions for estimating BD. This resulted in six different SOC stocks estimates from the following pedotransfer functions:

1. Saini (1966):  $BD = 1.62 - 0.06 * OM$ , Jeffrey (1970):  $BD = 1.482 - 0.6786 * (\log(OM))$ ,
2. Adams (1973):  $= 100 / (OM/0.244 + (100 - OM)/2.65)$ ,
3. Drew (1973):  $BD = 1 / (0.6268 + 0.0361 * OM)$ ,
4. Honeysett and Ratkowsky (1989):  $BD = 1 / (0.564 + 0.0556 * OM)$ ,
5. Grigal et al. (1989):  $BD = 0.669 + 0.941 * \exp(1)^{-0.06 * OM}$ . The  $OM = \text{organic matter content}$  was estimated as  $OM = \text{SOC concentration} * 1.724$ . These functions applied to the WoSIS data were selected because they were developed for a variety of soil weathering environments using multiple analytical techniques for measuring SOC and because they were recently proposed for the development of the United Nations global SOC map (Yigini et al., 2018). In addition, the WoSIS data have been curated under different protocols for controlling data quality and global interoperability standards (Batjes et al., 2017), and thus, this residual variance will allow us to have an idea of possible dispersion of values around the SOC calculated stocks.

### 2.6.2. Independent Datasets for Model Prediction

We calculated model residuals against two fully independent datasets across both countries ( $n=9239$ ). Across CONUS we used 6,179 SOC estimates (2010) from the Rapid Carbon Assessment Project (RaCA, Soil Survey Staff and Loecke, 2016; Wijewardane et al., 2016) and 3,060 (2009–2011) SOC estimates from top soil samples extracted from the Mexican National Forest and Soils Inventory of the Mexican Forest Service (2009–2011; Figure S2). These independent datasets have been collected using different sampling designs and using different SOC calculation methods from our initial training dataset (INEGI and ISCN).



The residual analysis against these independent datasets provides an overall measure of the models' sensitivity to multiple SOC data sources.

### 2.6.3. Spatial Autocorrelation of Model Residuals

We compared the spatial structure (i.e., spatial autocorrelation) of model residuals using linear geostatistics. The spatial structure accounts for the variance of values as a response of the geographical distance (e.g., meters) between SOC sampling points (Figure 2a). The spatial structure of a soil property can be quantified using variograms (a graphical method for modeling the relationship between distance between points and the variance of their values; Figure 2c) and the variogram parameters: nugget (uncorrelated variance), sill (spatially autocorrelated variance), and range (distance to the maximum variance) as explained previously (Oliver and Webster, 2014). We used automated variogram fitting (Hiemstra, et al., 2009) for calculating the nugget:sill ratio. The nugget is an uncorrelated component of soil variation that cannot be explained by our data; it depends on the calculation methods, the sampling resolution, and the spatial variability of SOC. The sill is the distance between the nugget and the variance stabilization ( $y$  axis) point while increasing distance ( $x$  axis). The range is the distance of the variance stabilization point. As in previous studies (Cruz-Cárdenas et al., 2014), a nugget:sill of  $<0.25$  was considered evidence of a strong spatial dependence, a relationship between 0.25 and 0.75 was considered a moderate spatial dependence, and a relationship  $> 0.75$  was considered a weak spatial dependence (Cambardella *et al.* 1994). We then used these variogram parameters to generate error maps by the means of Ordinary Kriging (Oliver and Webster, 2014), as explained earlier (Hengl et al., 2004), accounting for the potential spatially autocorrelation of the model residuals.

### 2.6.4. Model Residual Limits

For analyzing the model-based uncertainty, we estimated the quantile conditional response of the aforementioned modeling residuals to the best environmental covariates identified by our simulated annealing framework aiming to estimate model prediction limits. The main purpose of estimating model prediction limits is to identify the variance from the most probable predicted SOC stock for each pixel across the 250-m grids. For this purpose, we used the quantile regression forest approach, which is a variant of Random Forests. This method is able to (a) maintain the value of all observations in each node for each tree, not just their mean (as is the case of Random Forests), and (b) assesses the quantile conditional distribution at each predicted location (pixel). This method has the assumption that the full conditional estimated response is not different from the mean of the training dataset (Meinshausen, 2006). This method allowed us to quantify the maximum possible range of SOC prediction limits (e.g., 95%) given available data and available environmental covariates.

All analyses were performed in R (R Core Team 2018) and were repeated using subsets of available SOC data for the period 1991 to 2000 ( $n=4,877$ ) and for the period 2001–2010 ( $n=5,508$ ) in order to identify possible sensitivities of model predictions associated with defined (i.e., decadal) variations in training datasets.

## 3. Results

### 3.1. Descriptive Statistics

The harmonized SOC training dataset across both countries ( $n = 10,385$ , 1991–2010; Figure 2a) showed a right skewed distribution, and most estimates were between 0 and 10 kg/m<sup>2</sup> up to a maximum of 87.9 kg/m<sup>2</sup>. While the Mexican dataset dominates the 0- to 10-kg/m<sup>2</sup> range, the CONUS dataset has larger SOC values ( $>10$  kg/m<sup>2</sup>; Figure 2b). We used a logarithmic transformation (i.e.,  $\log(1+x)$ ) of the combined (Mexico-CONUS) dataset to reduce the skewed distribution for further analysis. The combined dataset shows a nugget:sill ratio of 0.55, suggesting a moderate spatial autocorrelation of its values, a sill of 0.9 and a nugget of 0.5 (units in  $\log(\text{kg m}^{-2} + 1)$ ; Figure 2c).

### 3.2. Recursive Feature Elimination

The five times repeated fivefold cross validation (applied to the recursive variable elimination framework) showed errors of 1.7 ( $R^2 = 0.30$ ), 2.0 ( $R^2 = 0.27$ ), and 2.6 ( $R^2 = 0.34$ ) kg/m<sup>2</sup> for the models 1991–2010 ( $n=10,385$ ), 1991 to 2000 ( $n=4,877$ ), and 2001–2010 ( $n=5,508$ ), respectively. For the years 1991–2010, the highest ranked environmental covariates for predicting SOC were the DEM and topographic terrain attributes (the wetness index, the valley bottom flatness index, and the valley depth index) and a remotely sensed precipitable water vapor estimate. For the years 1991–2000, the highest ranked environmental covariates were the land surface temperature, the DEM, the wetness index, the valley bottom flatness index, and the

standard deviation of the EVI (surrogate of vegetation seasonality). Finally, for the years 2001–2010, the highest ranked environmental covariates were: valley bottom flatness index, mean value of the EVI (surrogate of vegetation productivity), the valley depth index, the wetness index, and the night-time land surface temperature. These results showed consistency on the highest ranked environmental covariates such as the DEM and other terrain derivatives, considering the three recursive feature elimination models and years.

The same technique applied to independent datasets showed similar results (i.e., when combined RaCA and the Mexican Forest Service datasets). This independent analysis (years 2010–2012) showed an error of  $2.9 \text{ kg/m}^2$  ( $R^2=0.47$ ) using all environmental covariates and an error of  $3.4 \text{ kg/m}^2$  ( $R^2=0.33$ ) using just the highest ranked environmental covariates after the repeated cross validation. The highest five ranked environmental covariates of this model were the DEM and the topographic wetness index, the vegetation seasonality (standard deviation), and vegetation productivity (mean) from the EVI and mean monthly precipitation.

### 3.3. Simulated Annealing

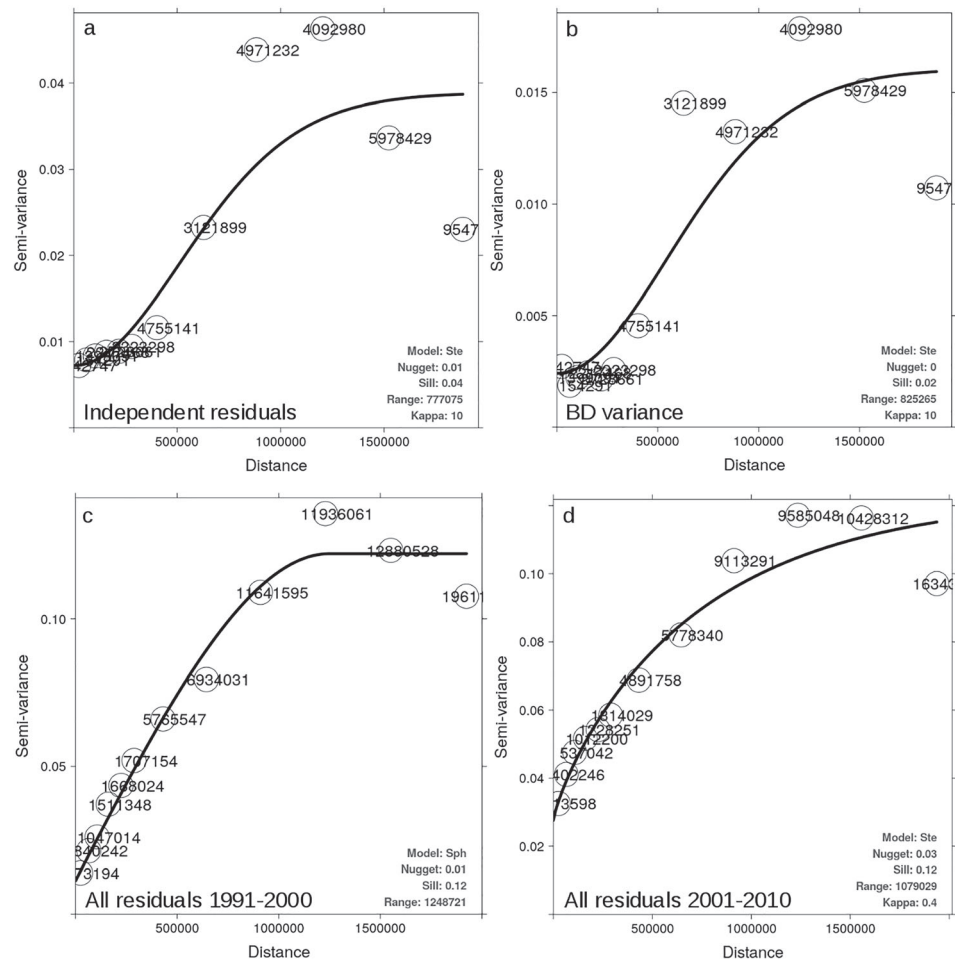
The simulated annealing framework confirmed the explanatory power of land surface temperature and precipitable water vapor, because these variables were consistently ranked as the highest environmental covariates in the three models (1991–2010, 1991–2000, and 2001–2010). For the three models, the simulated annealing framework revealed that mean annual precipitation and/or the total annual precipitation were also important predictors for the SOC dataset against a cross-validation strategy (5 times repeated, fivefold). The error estimates from this algorithm were similar compared with the previous analysis (see section 3.2), but with higher levels of explained variance for years 1991–2010 ( $2.2 \text{ kg/m}^2$ ;  $R^2=0.41$ ), years 1991–2000 ( $2.1 \text{ kg/m}^2$ ;  $R^2=0.31$ ), and years 2001–2010 ( $2.3 \text{ kg/m}^2$ ;  $R^2=0.46$ ).

The simulated annealing analysis on the independent datasets showed that a MODIS surface reflectance variable (M06MOD4; Table S1) becomes one of the first five important variables for predicting SOC. Other highest ranked environmental covariates for the independent datasets were also consistent with our previous results: the DEM, mean monthly precipitation, the standard deviation of the EVI, and precipitable water vapor. Modeling errors and explained variances ( $R^2$ ) on this model were also similar, with a mean error of  $3.1 \text{ kg/m}^2$  ( $R^2=0.42$ ) using only the highest ranked environmental covariates.

### 3.4. SOC Residual Analysis

We obtained six different SOC stocks (and mean errors) from the six pedotransfer functions used to estimate BD values (Figure S3). The equation provided by Drew (Drew, 1973:  $BD = 1/[0.6268 + 0.0361 * OM]$ ) was the best correlated function with our SOC prediction (1991–2010,  $r=0.4$ ). The residual variance of our predictions and the multiple SOC estimates derived from different BD pedotransfer functions had a standard deviation of 3.5, a median of 1.0, and mean variance of  $1.2 \text{ kg/m}^2$ . The residual variance of our predictions (1991–2010) against predictions from the independent model (RaCA- Mexican Forest Service;  $n=9239$ ) showed a standard deviation of 2.7, a median of 2.0, and a mean value of  $2.5 \text{ kg/m}^2$ . We report a moderate spatial structure (nugget:sill ratio of 0.25) for the residuals of our models and independent SOC estimates (Figure 4a). However, the residual variance of our predictions against the multiple SOC stocks calculated from different BD pedotransfer functions showed a strong spatial structure (nugget:sill ratio  $<0.1$ ) across both countries (Figure 4b).

When combining the residual variance from different BD pedotransfer functions and the residuals of the independent validation, we detected a significant increase of the nugget:sill ratio from 0.08 for the years 1991–2000 (Figure 4c) to a nugget:sill ratio of 0.25 for the period between 2001 and 2010 (Figure 4d). Thus, there was  $>100\%$  increase of uncorrelated spatial variation of SOC data (nugget:sill ratio increased from 0.08 to 0.25) from the model using 1991–2000 data to the model using 2001–2010 data. This increase of uncorrelated variation (increase of the nugget:sill ratio of  $>100\%$ ) was found in the combined residuals against independent data sets and against the BD pedotransfer function. These differences in the nugget:sill ratio are associated with inconsistencies in data sampling strategies and multiple collection periods of SOC data (Figure 4).

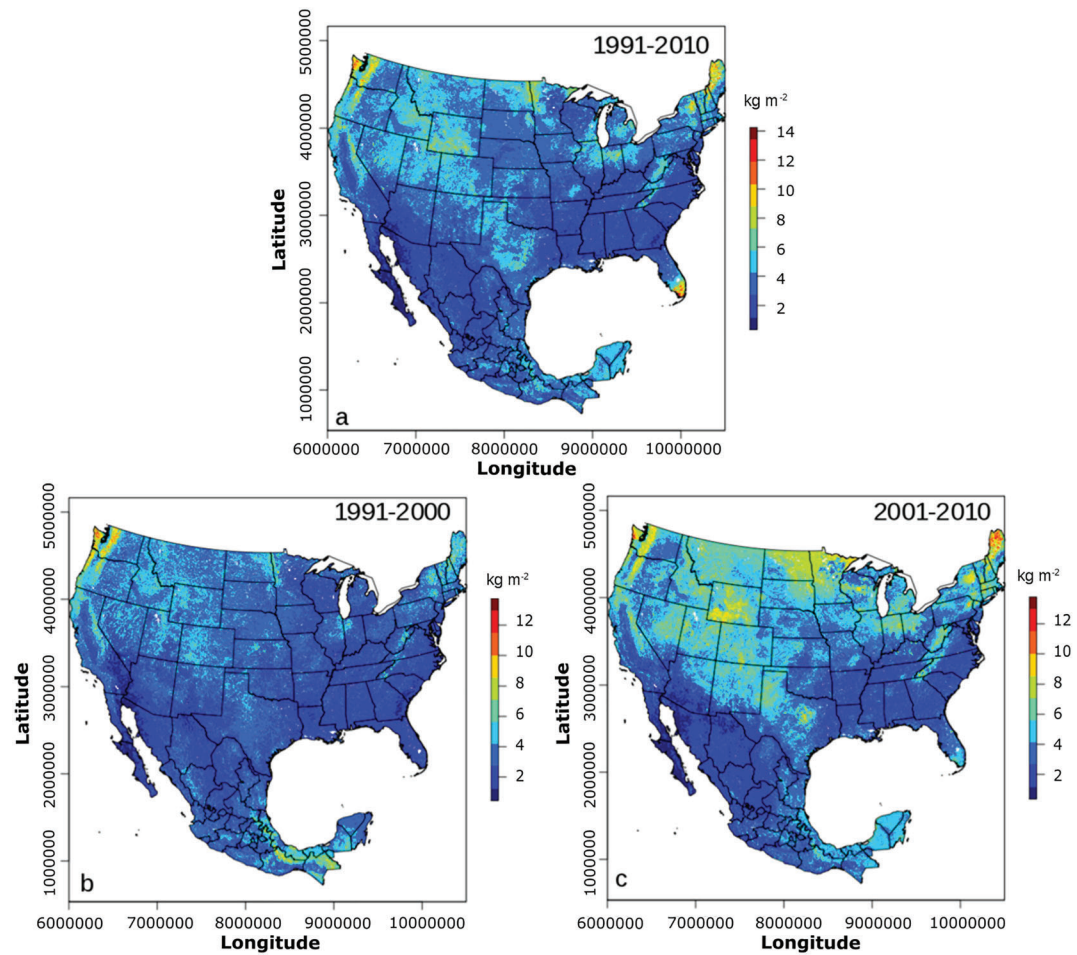


**Figure 4.** Variogram analysis applied to residuals of soil organic carbon (SOC) models. (a) The variogram of residuals against independent datasets. (b) The residual variance from the different pedotransfer functions used to calculate SOC stocks. The combined (independent models and pedotransfer functions) residuals for the periods (c) 1991–2000 and (d) 2001–2010. The numbers in the circles indicate the available pairs of points at a given distance. Variogram parameters are shown as insets on each plot: Sph = spherical model, Ste = Stein model parameterization (and its associated Kappa value).

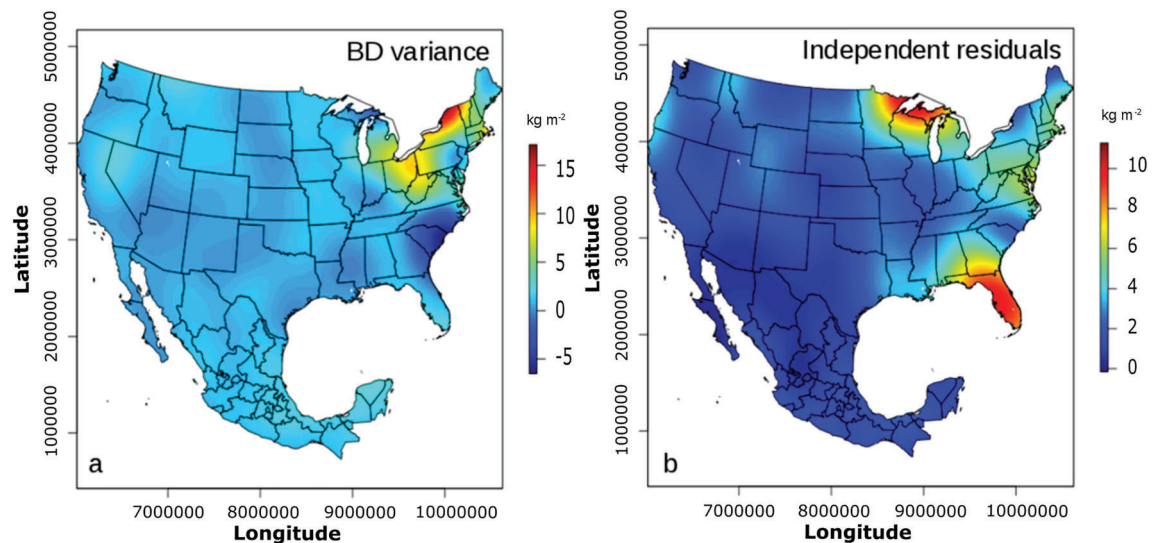
### 3.5. SOC Stocks

We estimated a total SOC stock (1991–2010) of 47 Pg (Figure 5a) that varies from 41 to 55 Pg of SOC for the models 1991–2000 (Figure 5b) and 2001–2010 (Figure 5c), respectively. For the years 1991–2010, the residual error map suggested  $10.4 \pm 5.1$  Pg of SOC variance associated with the use of multiple pedotransfer functions for BD and consequently calculating SOC stocks. The larger variance of associated with BD was found across the surroundings of the Great Lakes, in the states of Vermont, New York, and borders between Pennsylvania and Ohio, in CONUS (Figure 6a). The residual error map of our models against two fully independent datasets (RaCA and Mexican Forest Service) suggested a higher value of  $28.8 \pm 9.1$  Pg of SOC variance. The large variance associated with the independent datasets was found also across the surroundings of the Great Lakes, but in the states of Wisconsin and Minnesota (Figure 6b). Another large variance from the independent validation was found across the state of Florida, specifically across the south section in the everglades area where there are limited observations for the training dataset (Figure 1)

The estimated SOC stock after applying the same modeling strategy to the external datasets was 46 Pg of SOC (combined RaCA-Mexican Forest Service; Figure 7a), varying  $\pm 1$  Pg of SOC with the model 1991–2010 (ISCN plus INEGI, 1991–2010; Figure 5a). A linear model of our predictions against the independent datasets (RaCA-Mexican Forest Service) showed a mean error of  $1.0 \text{ kg/m}^2$  ( $R^2=0.43$ ; Figure S4).

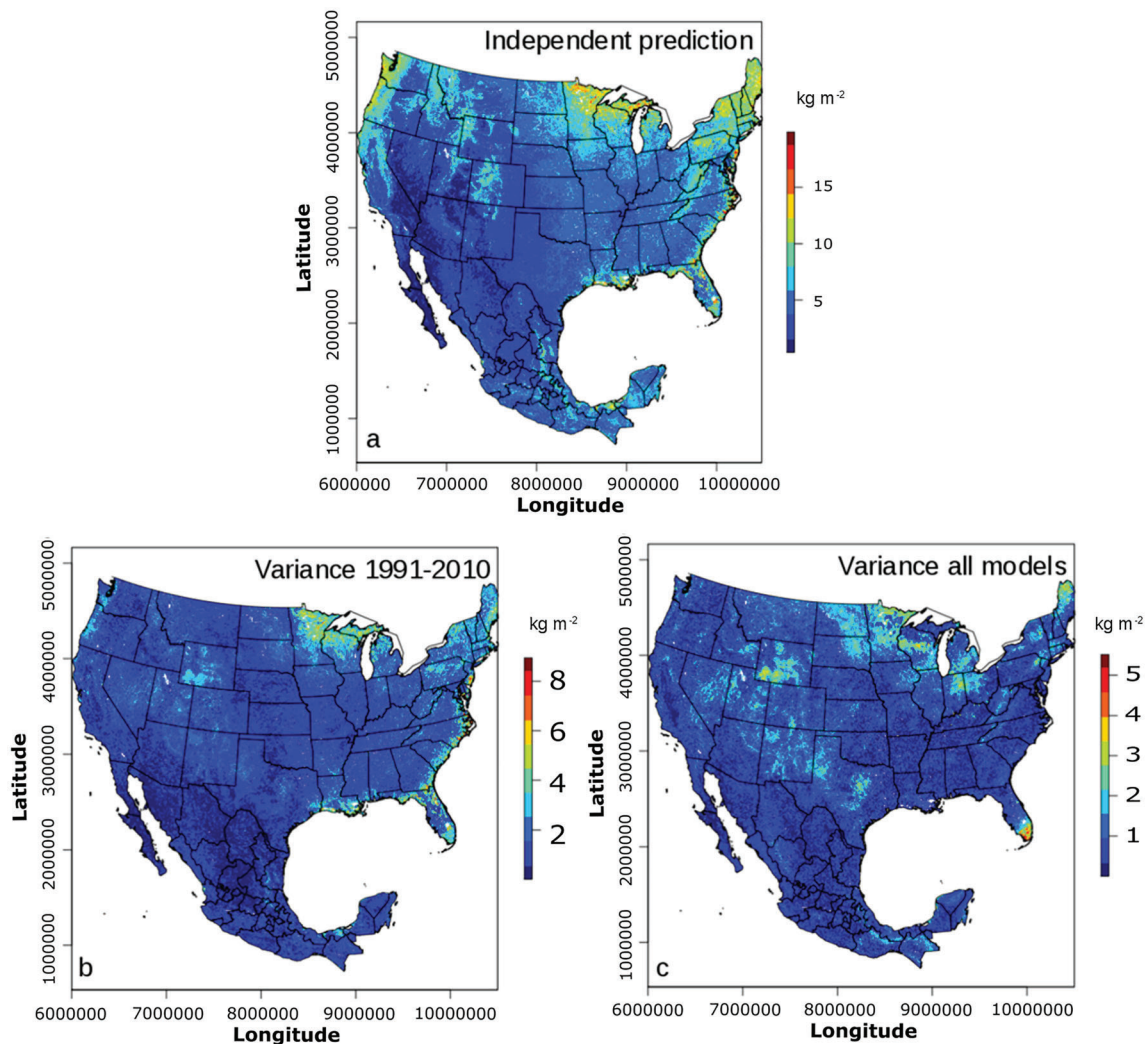


**Figure 5.** Predicted soil organic carbon across CONUS and Mexico. (a) Prediction using data for years 1991–2010. (b) Predictions with data for years 1991–2001. (c) Predictions with data for years 2001–2010.



**Figure 6.** Residual error maps interpolated using Ordinary Kriging. (a) The residual error map of pedotransfer bulk density (BD) residuals. (b) The residual error map of our models against independent validation datasets.





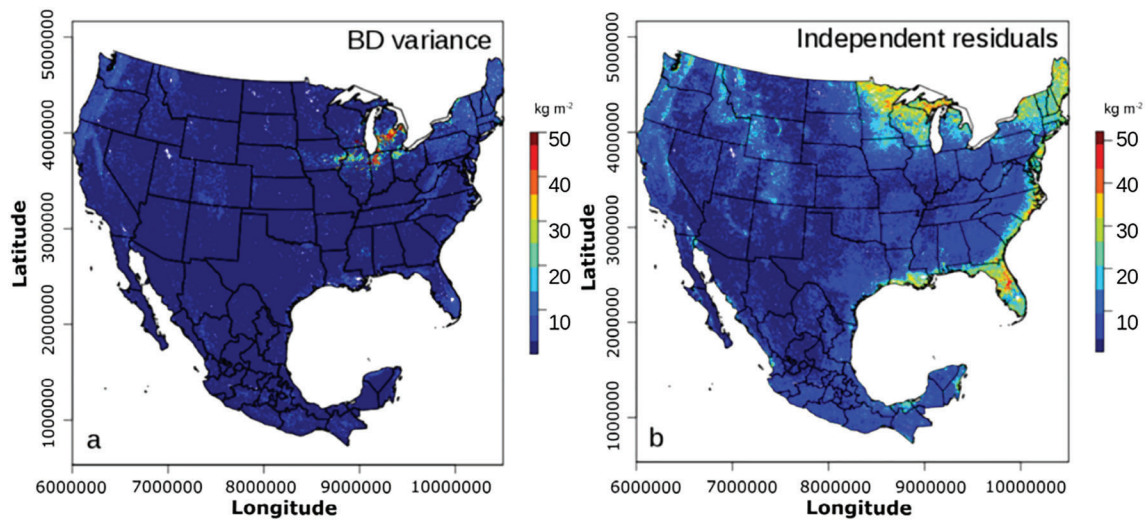
**Figure 7.** (a) Prediction of soil organic carbon generated using the independent datasets. (b) Model variance for predictions 1991–2000 and 2001–2010 using the Instituto Nacional de Estadística y Geografía (INEGI) and International Soil Carbon Network (ISCN) available data. (c) Variance of all SOC predictions (INEGI-ISCN, RaCA-Mexican Forest Service datasets).

The best correlation between SOC predictions was found between models 1991–2000 and 1991–2010 ( $r=0.8$ ) with  $\pm 5$  Pg of difference in the predicted SOC stocks. In contrast, the model for the year 2001–2010 was better correlated with the RaCA-Mexican Forest Service combined predictions ( $r=0.6$ ), but the SOC stocks varied for  $\pm 7$  Pg of SOC. The correlation between the model 1991–2010 and the independent analysis was lower ( $r=0.3$ ), but the SOC stocks showed less variation ( $\pm 1$  Pg of SOC).

The variance among all models based on INEGI and ISCN data was  $\pm 7.6$  Pg of SOC (Figure 7b). This value increased up to  $\pm 12$  Pg of SOC by adding the variance of the models based on the independent analysis (Figure 7c). Thus, we provide a SOC stock for both countries between 46 and 47 Pg of SOC with a total modeling variance of  $\pm 12$  Pg.

### 3.6. Quantile Conditional Distribution of Residuals

The quantile conditional distribution (used to identify the model prediction limits) for the residuals against fully independent datasets suggest a maximum possible SOC variance of  $\pm 73$  Pg of SOC. This is the SOC variance from the full quantile conditional response of these residuals to the highest ranked environmental covariates. From the BD pedotransfer function variance, the full conditional response to the highest ranked environmental covariates showed a lower value of  $\pm 20$  Pg of SOC. Thus, less uncertainty was found from



**Figure 8.** Conditional quantile distribution of soil organic carbon residuals to the highest ranked environmental covariates from (a) the bulk density (BD) residual variance and for (b) the residual variance against models generated with fully independent datasets.

the use of multiple pedotransfer functions than from validating against fully independent datasets. These results highlight the large variance of possible SOC predictions given the use of multiple training data sources constrained to a relatively short (i.e., two decade) period of time. The model-based uncertainty results are shown in Figure 8. For the BD pedotransfer function variance, the larger range of model based uncertainty was found across the Great Lakes of northern CONUS and the border with Canada (Figure 8a), The model-based uncertainty using independent residuals shows the largest values across, Florida, the east coast and the surroundings of the Great Lakes in CONUS, and some areas of southeast Mexico (Figure 8b).

#### 4. Discussion

We predicted SOC (across Mexico and CONUS at 0–30 cm of soil depth) and generated gridded estimates at 250-m spatial resolution for the period 1991–2010. We provided predictions of SOC using multiple inputs of data and a feature selection framework (recursive elimination of predictors and the simulated annealing algorithms) that allowed us to identify the most informative SOC environmental covariates to determine SOC stocks between 1991 and 2010. We calculated SOC stocks for both countries ( $46\text{--}47 \pm 12$  Pg), and these values were  $>30\%$  below previous global estimates such as the SoilGrids system or the Harmonized World Soil Database. We have highlighted large discrepancies between modeling outputs based on multiple data collection periods (Figure 5) and between global SOC products such as the state of the art SoilGrids250m (Figure S5). Furthermore, our results have implications for the use and interpretation of SOC legacy data or aggregated SOC information. Specifically, we found a large difference for predicted SOC stocks (from 41 to 55 Pg of SOC) between 1991–2000 and 2001–2010 that cannot be fully attributed to SOC dynamics but also to inconsistencies in the spatial configuration of available datasets, the use of different SOC calculation methods, and the different periods of data collection. These results open new research questions about the interpretation of apparent changes in SOC stocks across time, and future studies should determine (a) if regional-to-global differences are due to active management practices/land use change or (b) if apparent changes are overshadowed by large uncertainty estimates due to inconsistencies in methods and modeling variability.

##### 4.1. Highest Ranked Environmental Predictors

Our results suggest that using a few informative environmental predictors (e.g., the DEM and terrain derivatives, the EVI, or precipitation patterns) for the spatial variability of SOC have a similar performance ( $\sim 50\%$  of explained variance) to a high-dimensional covariate space (the 150 environmental covariates reported in Table S1) across Mexico and CONUS and using the available datasets between 1991 and 2010.

The use of a high-dimensional covariate space to predict SOC (and other soil properties) may be needed to maximize prediction accuracy at the global scale (Hengl et al., 2017); however, for local to regional applications, some environmental covariates for SOC may be statistically redundant and lead to unnecessarily increases of the computing resources required for prediction purposes. Reducing the statistical redundancy of environmental covariates in SOC models across multiple spatial resolutions will simplify computing requirements and model complexity (Guo et al., 2019). Reducing the complexity of SOC models would be appealing in further applications of SOC spatial information (e.g., land carbon uptake modeling, climate system modeling, and niche modeling), which require similar predictors as for SOC.

Our simulated annealing analysis highlights SOC relationships (positive and negative) with climate variables (precipitation and land surface temperature), elevation (and terrain derivatives), and vegetation greenness (productivity and seasonality) that are consistent with previous literature describing SOC drivers across diverse environmental conditions (Evans et al., 2011; Hobbey et al., 2015). In addition, when applying the simulated annealing framework to the independent datasets, a MODIS surface reflectance short wave infrared band (M06MOD4; Table S1) was one of the first five important variables predicting SOC, which is consistent with the infrared based methods used by the USDA for developing of the RaCA dataset (Wijewardane et al., 2016). The prediction capability of the infrared spectra (e.g., near infrared and midinfrared) for SOC can be attributed to the strong spectral absorption characteristics of soil organic matter and BD (the main components of SOC) in the infrared spectral bands (Guo et al., 2019). Thus, the main relationships driving our SOC predictions can be interpreted and associated with the use of different data inputs and specific environmental covariates (e.g., the DEM and terrain parameters, the EVI, and the MODIS surface reflectance infrared data, precipitation, and temperature gridded surfaces) that can be periodically acquired from remote sensing at the global scale.

## 4.2. Uncertainty Quantification

### 4.2.1. BD Pedotransfer Functions

We report ~10Pg of SOC variance associated with SOC calculation inputs. The combined and quality-controlled dataset we used have been processed following international standards for increasing precision and accuracy (Batjes et al., 2017; Harden et al., 2017). However, the major limitation of these datasets is arguably the low availability of BD and CF data. Inconsistencies in BD and CF data could explain the large variance found (Figure 6a) across the highly productive landscapes of the north east of CONUS. It has been discussed that SOC stocks are systematically overestimated by misuse of the BD and CF content parameters (Poeplau et al., 2017). Thus, correction of BD is fundamental to achieve realistic SOC estimates and to reduce the potential overestimation of SOC stocks (Köchy et al., 2015). The lack of accurate BD and CF data and the large variance of the global SOC values are key issues that could explain the discrepancy between country-to-global SOC estimates (Tifafi et al., 2017). Thus, our results provide a spatially explicit measure of SOC variance derived from six conventional BD pedotransfer functions that can be used to explain discrepancies between national, regional, and global SOC estimates.

### 4.2.2. Spatial and Temporal Variations of Available Data

We found differences in the spatial structure (i.e., autocorrelation) of modeling residuals from multiple models and periods of time (1991–2000 and 2001–2010), that result in large differences on predicted stocks (from 41 to 55 Pg of SOC) from these defined periods of time. This period of time (1991–2010) have experienced intensive land use and environmental changes across both countries, and our results could be used for identifying sensitive areas of SOC changes or areas that require further research (Figure 5). However, a previous study suggested that SOC could increase under reforestation conditions around 2 Pg per century in the topsoil (Nave et al., 2018), so our “decadal” modeling results may be overestimating the SOC gain between those time-periods. Here, we discuss our results under this consideration.

While we detected a reduction of SOC across most of Mexico, we detected a larger gain of SOC mainly across higher latitudes of CONUS (when comparing models between 1991–2000 and 2001–2010). Recent efforts have shown multiple agricultural practices that can lead to substantial SOC gains (Singh et al., 2018), and SOC gains have been reported on previous studies across higher latitudes of CONUS in response to agricultural practices (Adhikari & Hartemink, 2017). For example, alpine forests have been recognized as important SOC sinks under warming conditions (Ding et al., 2017). Recent reports have shown that some land carbon uptake models tend to project increases in high-latitude SOC that are inconsistent with empirical

studies that indicate significant losses of SOC with predicted climate change across these areas (Lajtha et al., 2018). The uncertainty of current SOC available information is one limiting factor for increasing the agreement and explaining the aforementioned inconsistencies of SOC models (Crowther et al., 2016). When applying the analysis independently on the specific decades (1991–2000 and 2001–2010), we were forced to remove large amounts of data across large geographical areas; consequently, these areas were not equally represented (in terms of data information) on these models. We argue that the spatial distribution and statistical differences on data available for SOC models can explain discrepancies of SOC trends, as previously shown at the global scale (van Gestel et al., 2018). Reducing the amount of training data increases modeling errors and the uncertainty of SOC predictions (Lagacherie et al., 2019). Thus, we highlight that caution must be taken when limited amount of information is used to predict SOC stocks across large geographical areas and then use that information to quantify apparent changes in SOC stocks without considering uncertainty. In this study, rather than reporting a SOC change between decades, we postulate that a better practice is to use all available data (1991–2010) to increase spatial representation. Thus, we were able to model SOC spatial variability and compared results with two fully independent datasets to determine the most probable SOC stock estimate (46 to 47 Pg of SOC) for the period around 1991 and 2010.

#### 4.2.3. Quantile Response of Residual Variance

Model prediction limits from the full quantile response of independent model residuals indicated a larger SOC variance across both countries (up to 73 Pg of SOC variance) than the full response of the residual variance associated with the BD pedotransfer functions (20 Pg of SOC). These results are useful for benchmarking SOC models and represent a valuable complement for the uncertainty indicators of the predicted SOC spatial variability (Lagacherie et al., 2019). This variance relies on a nonparametric and accurate way of estimating conditional quantiles and the overall reliability of tree-based ensembles such as Random Forests (Meinshausen, 2006). This approach has been used for analyzing the uncertainty on soil mapping applications, and larger uncertainties have been reported when reducing the data availability in numerical experiments (Lagacherie et al., 2019; Vaysse & Lagacherie, 2017). Thus, we provide multiple uncertainty indicators as they are useful to better interpret model limitations associated with available datasets and complement (across unsampled areas) our cross validation and independent validation results. We propose that the results of this quantile analysis applied to SOC modeling residuals could be used for identifying areas that require higher sampling effort due larger discrepancies of multiple SOC model predictions.

#### 4.3. SOC Stocks Across CONUS and Mexico

The estimated SOC stock across both countries (46–47 Pg of SOC) could be used for quantifying the contribution of SOC to the regional (e.g., North America) carbon cycling for the analyzed period of time (1991–2010). Our predicted SOC stock is lower when compared to values obtained from global estimates such as the re-gridded HWSD (Wieder et al. 2014) or the SoilGrids250m system (Hengl et al., 2017), where this value increases to ~71 and ~92 Pg of SOC, respectively. High discrepancy of these two global products has been reported earlier at global- (Tifafi et al., 2017), country-, or region-specific scales (Guevara et al., 2018, Vitharana et al., 2019). Moreover, our results showed discrepancies comparing country-specific studies reporting SOC stocks in CONUS (29.3 Pg of SOC; Bliss et al., 2014) and Mexico (9.15 Pg; Cruz-Gaistardo & Paz-Pellat, 2014, Paz Pellat et al., 2016), as we report ~39 Pg of SOC for CONUS and ~7 Pg of SOC for Mexico in the first 30 cm of soil. Our results highlight the need to provide country-to-region specific estimates using the best available datasets, to improve global SOC estimates by developing analytical frameworks for optimizing multiple SOC modeling efforts and sampling strategies (Guevara et al., 2018).

We report a density of SOC across CONUS (4.98 kg/m<sup>2</sup>) that was relatively higher than the soils of Mexico (4.22 kg/m<sup>2</sup>). Globally, the soil carbon pool (at 1-m depth) is estimated to have around 1,500–2,400 Pg (Paustian et al., 2019), while the SOC pool in the upper 30 cm is estimated to be  $755 \pm 119$  Pg (Batjes, 2016). Our results suggest that Mexico represents ~1% and CONUS ~5% of the global SOC pool at 30-cm depth. Recent revisions highlight that the SOC stock at 30-cm soil depth remains unclear (Lajtha et al., 2018), and this study provides new insights for interpreting the discrepancies around the topsoil SOC pool across CONUS and Mexico.

Our SOC estimates across forested areas are comparable to those reported on studies (Bolaños González et al., 2017; Domke et al., 2017). However, our results show high uncertainty across areas dominated with high SOC values (e.g., >1 g/cm<sup>2</sup>, some northern and tropical forests, peatlands, and other black soils



dominated areas) and across higher latitudes (Figure 6), as documented in previous studies (Tian et al., 2015). Unfortunately, these areas are poorly represented in the available datasets (<10% of available data with values  $>1 \text{ g/cm}^2$ ) and we encourage future monitoring efforts to increase their representativeness.

#### 4.3.1. SOC Across Land Cover Classes

Our study confirms the presence of important SOC stocks across both forest and agricultural soils. Across both countries, we found higher SOC stocks in croplands, representing 26% of total SOC within the upper 30cm of soil. Across Mexico, we found that 42% of SOC was stored in forest soils and 24% in agricultural soils, while 31% of SOC across CONUS is stored in forest soils and 27% in agricultural soils. While organic matter-rich and deep soils dominate most agricultural areas across CONUS (Adhikari & Hartemink, 2017), most agricultural soils in Mexico tend to be shallow (Guerrero et al., 2014, ~30cm depth); consequently, we emphasize that carbon management, monitoring, and conservation strategies must be developed from a country-specific approach considering country-specific land cover classes.

We found that tropical or subtropical broadleaf evergreen forests are the natural vegetation class with the highest SOC pool across Mexico (1.22 Pg), while temperate broadleaf deciduous forests had the highest SOC pool across CONUS (6.41 Pg). Grasslands and shrublands are also important SOC reservoirs, as they store around 37.7% of SOC across Mexico and 34.9% of SOC across CONUS (Table S2). Such estimates are relevant for public policy around SOC conservation efforts (e.g., FAO, 2017) because grasslands and shrublands transitions are increasingly vulnerable to global warming and the increase of aridity conditions, which would result in a decrease of SOC stocks (FAO, 2017). Thus, accurately quantifying the spatial variability of SOC across grasslands and shrublands will be an important component for enhancing SOC sequestration by better informing conservation efforts of soil ecosystem functions across North America.

Accurate SOC estimates represent a key variable to quantify human induced disturbances to the carbon cycle across land cover classes. We report that temperate forests of CONUS contain the larger SOC reserves, while tropical or subtropical broadleaf evergreen forest and wetlands are the land cover classes with higher SOC in Mexico than CONUS (Table S2). Respectively, the tropical or subtropical broadleaf evergreen forests are the most productive ecosystems of Mexico (Murray-Tortarolo et al., 2016). The wetland category, with high carbon sequestration potential, includes mangroves (and other coastal wetlands), which have been recognized as the ecosystems with higher carbon storage capacity from the site-specific to the global scales (Adame et al., 2015; Atwood et al. 2017; Vázquez-Lule et al., 2019). Our results represent benchmarks for SOC monitoring across these land cover classes. Thus, the spatial predictions of SOC at 250m allows for the interpretation of SOC spatial patterns across land cover classes of Mexico and CONUS accounting for sensitivities associated with the use of multiple data inputs.

#### 4.4. Final Remarks

Optimizing future SOC sampling strategies while reducing modeling variance and increasing model agreement against model independent datasets collected under different circumstances (e.g., logistics, design, main purpose, and SOC estimation methods) are large challenges for enabling SOC carbon mapping and monitoring systems. New and better SOC parameters are required for reducing the current discrepancy between multiple sources of SOC data (Guevara et al., 2018; Tifafi et al., 2017) and enabling SOC monitoring systems (Viscarra-Rossel et al., 2014). The lack of accurate SOC spatial information and the combination of multiple SOC data sources could result in large variance estimates across the two countries (e.g., red areas of Figure 6). We propose that areas with high variance suggest that these regions require higher SOC sampling efforts.

We provide high spatial resolution (e.g., 250-m pixels) SOC estimates that account for model uncertainty. Such estimates are needed for identifying regions that should be targets for SOC protection (Lagacherie & McBratney, 2006). Soil carbon protection is increasingly important to restore the current negative imbalance in our soil carbon budget due, for example, to the development of agricultural systems and croplands (Sanderman et al., 2017). Accurate SOC estimates at the relevant (local) scale for farmers and landowners (e.g., spatial resolution of 250m or less) would be an important component to reduce land degradation and improve the efficiency of current efforts for sequestering SOC (Bonfatti et al., 2016; Malone et al., 2017). Thus, our results provide insights for identifying and delineating land areas with high potential for SOC stocks that account for model sensitivity to multiple data inputs and sources. Finally, this research is

timely because there is high discrepancy between SOC global estimates that needs to be solved in order to better quantify SOC dynamics (Tifafi et al., 2017). Consequently, this discrepancy can influence the estimates of SOC warming response (Karhu et al., 2010) and the carbon-climate feedback that could accelerate climate change (Crowther et al., 2016). We hope that this study motivates an increase in country-specific soil surveys, data sharing, and modeling of SOC estimates at higher spatial resolution with a better quantification of uncertainty.

## Data Availability Statement

All modeling output is available through The Oak Ridge National Laboratory Distributed Active Archive Center (ORNL DAAC) for biogeochemical dynamics of the NASA-Earth Observing System Data and Information System (DOI: <https://doi.org/10.3334/ORNLDAAC/1737>). Public data to parameterize the model are available from the Rapid Carbon Assessment Project ([https://www.nrcs.usda.gov/wps/portal/nrcs/detail/soils/survey/?cid=nrcs142p2\\_054164](https://www.nrcs.usda.gov/wps/portal/nrcs/detail/soils/survey/?cid=nrcs142p2_054164)), the International Soil Carbon Network (<https://iscn.fluxdata.org/>), the Instituto Nacional de Estadística y Geografía (INEGI, SERIES 1 and 2;  $n > 65,000$  pedons available; Krasilnikov et al., 2013), and covariates from the SoilGrids250m project ([https://soilgrids.org/#!/layer=ORCDRC\\_M\\_sl2\\_250m&vector=1](https://soilgrids.org/#!/layer=ORCDRC_M_sl2_250m&vector=1); Hengl et al., 2017).

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